

Hierarchical modeling of endpoints of different types with generalized linear mixed models

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Introduction

- In many situations, one is confronted with the analysis of longitudinal data
- Often, measurements (or sequences) of both continuous and categorical outcomes are recorded
- Multivariate methods for continuous outcomes are well understood
- Methods for joint continuous and discrete outcomes are less familiar
- Complexity of joint model increases with number of outcomes

Why Jointly?

- Association structure can be of interest
 - E.g. Relationship between outcome-specific evolutions
- Interest in comparison of trends for different outcomes
 - E.g. testing difference in treatment effect between many outcomes
- E.g. jointly testing treatment effect on a set of outcomes
- Discriminant analysis, principal component analysis

Bivariate Model for Binary and Continuous Longitudinal Outcomes

We want a method that:

- Allows to specify the marginal model for each outcome
- Estimate the bivariate association between continuous and discrete outcomes
- Accounts for the longitudinal effects / longitudinal association
- Is computationally not too complex

Factorization Models

- Models that condition on the continuous outcome

$$f(C, B) = f(C) \times f(B|C)$$

- Models that condition on the binary outcome

$$f(C, B) = f(B) \times f(C|B)$$

- Correlation difficult to characterize with conditional joint models

- Difficult to generalize to three or more endpoints

E.g. Tate (1954), Olkin and Tate (1961), Little and Schluchter (1985), Krzanowski (1988)

E.g. Catalano and Ryan (1992), Fitzmaurice and Laird (1993)

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Background		
	<p>Direct Specification of Joint Model</p> <ul style="list-style-type: none"> • Probit Approach <ul style="list-style-type: none"> ◦ assume normal latent variable ◦ use correlation coefficient • Copula function <ul style="list-style-type: none"> ◦ e.g. Plackett latent variable using odds ratio <p>GLMM Approach</p> <ul style="list-style-type: none"> • allow the link function to change with the nature of the outcomes • use random effects or residual error to account for the association in the data <p>Burzykowski et al., 2001; Geys et al., 2001; Faes et al., 2004; de Leon and Wu, 2011 Faes et al. (2006), Gueorguieva and Sanacora (2006), Regan and Catalano (1999)</p>	

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A General Linear Mixed Model		
	$Y_i = \boldsymbol{\mu}_i + \epsilon_i = h(\mathbf{X}_i \boldsymbol{\beta} + \mathbf{Z}_i \mathbf{b}_i) + \epsilon_i$ <p>with</p> $\mathbf{b}_i \sim N(\mathbf{0}, \mathbf{D})$ <ul style="list-style-type: none"> • $\mathbf{Y}_i = (\mathbf{C}'_i, \mathbf{B}'_i)'$ • $h(\cdot)$ is allowed to change with nature of outcome <ul style="list-style-type: none"> ◦ Identity link for continuous component ◦ Logit link for binary component • ϵ_i is the residual error structure, of which the variance depends on the mean-variance links of the various endpoints 	

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Correlated random effects model

Assume the following bivariate model:

$$\begin{pmatrix} Y_{i1j} \\ Y_{i2j} \end{pmatrix} = \begin{pmatrix} \alpha_0 + \alpha_1 X_{ij} + b_{i1} \\ \exp(\beta_0 + \beta_1 X_{ij} + b_{i2}) \end{pmatrix} + \begin{pmatrix} \epsilon_{i1j} \\ \epsilon_{i2j} \end{pmatrix}$$

where the random effect b_{i1} and b_{i2} are normally distributed as

$$\begin{pmatrix} b_{i1} \\ b_{i2} \end{pmatrix} \sim N \left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \tau_1^2 & \rho\tau_1\tau_2 \\ \rho\tau_1\tau_2 & \tau_2^2 \end{pmatrix} \right)$$

and where ϵ_{i1j} and ϵ_{i2j} are independent

Remark: Standard software can be used to obtain parameter estimated for this bivariate model (e.g., by integrating out random effects using SAS-procedure NL MIXED)

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Bivariate Continuous/Binary Model

- The correlation among the continuous and binary endpoints is approximately equal to:

$$\rho_{Y_{i1j} Y_{i2j}} \approx \rho \frac{\tau_1}{\sqrt{\tau_1^2 + \sigma^2}} \frac{v_{i2j}\tau_2}{\sqrt{v_{i2j}^2\tau_2^2 + v_{i2j}}}.$$

- The correlation among the continuous endpoints:

$$\rho_{Y_{i1j} Y_{i1k}} \approx \frac{\tau_1^2}{\tau_1^2 + \sigma^2}.$$

- The correlation among the binary endpoints:

$$\rho_{Y_{i2j} Y_{i2k}} \approx \frac{v_{i2j}^2\tau_2^2}{v_{i2j}^2\tau_2^2 + v_{i2j}}.$$

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Bivariate GLMM Model	

In summary:

- Marginal distribution are specified, depending on type of endpoint
- Random effect used to account for longitudinal effect
- Correlation between the random effects used to account for association among the endpoints
- Can be used for any mixture of endpoints:
 - Binary/Continuous
 - Binary/Binary
 - Continuous/Continuous
- Not restricted to a bivariate model

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Example: Repeated-Dose Toxicity Study (Irwin's study)

Research Question:

Has the chemical an effect on general activity and behavior?

- Male rats were dosed during 3 consecutive days
- Randomized in control (5 rats) and treatment group (15 rats)
- Examination of rats at 2, 4, 6, 8 and 24 hours after daily exposure
- A set of observational and interactive measurements are used (Irwin's method)

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Repeated-Dose Toxicity Study

Variable	Type	Description	Classification
1 Locomotor Activity	Binary	Characterized by abnormal biting, restlessness, writhing	Behavioral
2 Pinna Reflex	Binary	Animal's twitch reflex upon being touched on the auricle	Spontaneous activity
3 Toe Pinch	Binary	Animal's response to pain upon having the toe squeezed	Sensoro-motor response
4 Positional Passivity	Binary	Animal's struggle response to sequential handling	Behavioral
5 Grip Strength	Continuous	Animal's forelimb grip strength	Neurologic
6 Pupil Size	Continuous	Animal's pupil diameter	Muscle tone
7 Temperature	Continuous	Animal's body temperature	Autonomic
8 Vocalization	Continuous	Animal's incidence of squeaking during manipulation	Autonomic
			Motor-affective response

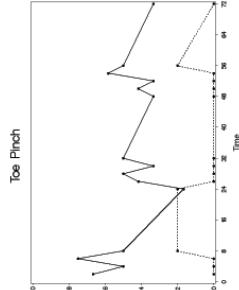
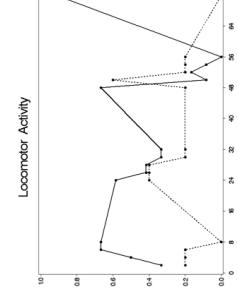
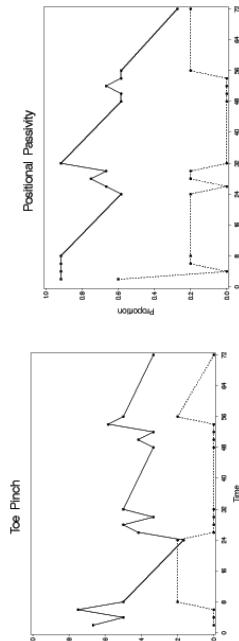
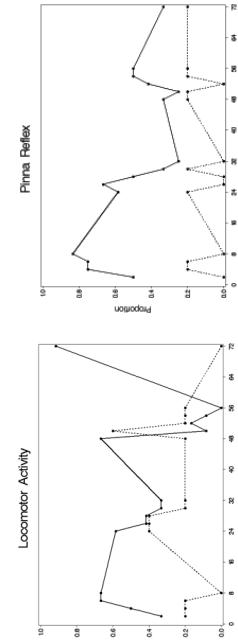
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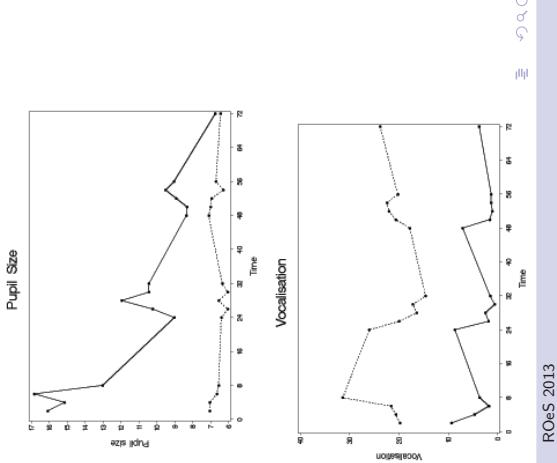
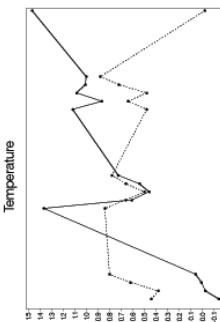
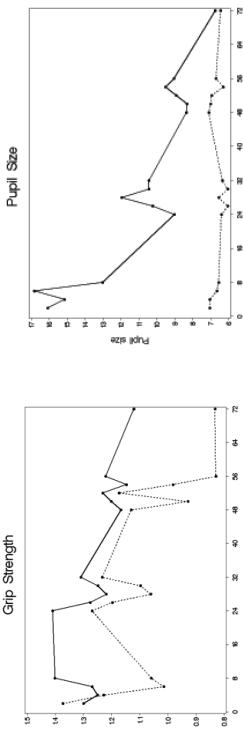
Four Binary Endpoints



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Four Continuous Endpoints



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Application: Joint Model

$$h_k^{-1}(\mu_{ij}) = \beta_{0k} + \beta_{1k}g_i + \beta_{2k}t_{ij} + \beta_{3k}d_{ij} + \beta_{4k}t_{ij}d_{ij} + \beta_{5k}g_i t_{ij} + \beta_{6k}g_i d_{ij} + b_{ik},$$

with

$$\mathbf{b}_i = \begin{pmatrix} b_{i1} \\ b_{i2} \\ \vdots \\ b_{i8} \end{pmatrix} \sim N \left\{ \begin{pmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{pmatrix}, \begin{pmatrix} \tau_1^2 & \rho_{12}\tau_1\tau_2 & \cdots & \rho_{18}\tau_1\tau_8 \\ \rho_{12}\tau_1\tau_2 & \tau_2^2 & \cdots & \rho_{28}\tau_2\tau_8 \\ \vdots & \ddots & \cdots & \vdots \\ \rho_{18}\tau_1\tau_8 & \rho_{28}\tau_2\tau_8 & \cdots & \tau_8^2 \end{pmatrix} \right\}.$$

- h_k is the identity link in case of a continuous endpoint and the logit link in case of a binary endpoint
- g_i is indicator variable for treatment
- t_{ij} is the time after exposure
- d_{ij} is the day of the experiment

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Application: Joint Model

$$h_k^{-1}(\mu_{ij}) = \beta_{0k} + \beta_{1k}g_i + \beta_{2k}t_{ij} + \beta_{3k}d_{ij} + \beta_{4k}t_{ij}d_{ij} + \beta_{5k}g_it_{ij} + \beta_{6k}g_id_{ij} + b_{ik},$$

with

$$\mathbf{b}_i = \begin{pmatrix} b_{i1} \\ b_{i2} \\ \vdots \\ b_{i8} \end{pmatrix} \sim N \left\{ \begin{pmatrix} 0 & & & \\ 0 & \tau_1^2 & \rho_{12}\tau_1\tau_2 & \cdots & \rho_{18}\tau_1\tau_8 \\ \vdots & \rho_{12}\tau_1\tau_2 & \tau_2^2 & \cdots & \rho_{28}\tau_2\tau_8 \\ 0 & \cdots & \cdots & \cdots & \vdots \\ \rho_{18}\tau_1\tau_8 & \rho_{28}\tau_2\tau_8 & \cdots & \tau_8^2 & \end{pmatrix} \right\}.$$

- h_k is the identity link in case of a continuous endpoint and the logit link in case of a binary endpoint
- g_i is indicator variable for treatment
- t_{ij} is the time after exposure
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Problem

No convergence!

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Problem with Extension to Higher Dimensions

The marginal likelihood contribution for subject i becomes

$$\ell_i(\Theta | \mathbf{Y}_{i1}, \dots, \mathbf{Y}_{im}) = \int_{\mathbb{R}^{m^n}} \prod_{j=1}^n f_{ij}(y_{i1j}, \dots, y_{imj} | b_i, \Theta) f(\mathbf{b}_i | D) d\mathbf{b}_i.$$

Computational problems often arise when m increases due to the m -dimensional integral, especially when outcomes are of different type.

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Pseudo-Likelihood (Composite Likelihood)		

Replace the full likelihood contribution for subject i

$$\ell_i(\Theta | \mathbf{Y}_{i1}, \dots, \mathbf{Y}_{im})$$

by the pseudo-likelihood function

$$p\ell_i = \prod_{k=1}^{m-1} \prod_{l=k+1}^m \ell_{ikl}(\Theta | \mathbf{Y}_{ik}, \mathbf{Y}_{il}),$$

where each contribution ℓ_{ikl} is the likelihood corresponding with a bivariate GLMM model for outcomes k and l .

Remark

Standard software can be used to obtain parameter estimated for this model (e.g., SAS-procedure NLMIXED).

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Idea Pseudo-Likelihood		

- Replace full likelihood by function that is easier to evaluate
- Estimates all pairwise correlations
- Use of sandwich variance estimator to adjust for potential misspecification
- Pseudo-likelihood estimates are consistent and asymptotically normal
- Loss of efficiency is small for realistic settings
- Pseudo-likelihood ratio test statistic easy-to-compute

Arnold and Strauss (1991)

Geyls, Molenberghs and Ryan (1999)

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Idea Pseudo-Likelihood		

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Pairwise Approach	
Pseudo-likelihood Approach	$\ell_i(\Theta \mathbf{Y}_{i1}, \dots, \mathbf{Y}_{im})$
Pairwise Approach	$p\ell_i = \prod_{k=1}^{m-1} \prod_{l=k+1}^m \ell_{ikl}(\Theta \mathbf{Y}_{ik}, \mathbf{Y}_{il})$
Pseudo-likelihood Approach	$p\ell_i = \prod_{k=1}^{m-1} \prod_{l=k+1}^m \ell_{ikl}(\Theta_{k,l} \mathbf{Y}_{ik}, \mathbf{Y}_{il})$
	<ul style="list-style-type: none"> maximize the full pseudo-likelihood function at once number of parameters to estimate jointly increases with m pseudo-likelihood ratio tests can be used
	Faes et al. (2007)
	<ul style="list-style-type: none"> maximize all pairwise likelihood functions separately number of pairwise models increases with m estimates of joint model derived by taking averages over estimates from the pairwise models
	Fieuws and Verbeke (2005)
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Model Fitting

- Pseudo-likelihood and pairwise method used to estimate parameters
 - Relatively easy to program
 - Parameter estimates from the two approaches are very similar

Estimated Correlations

Estimated correlation matrix for random effects

Values on the diagonal are the intra-class correlations corresponding to each outcome.

Represents relationship among the subject-specific deviations from different outcomes.

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Estimated Correlations

Estimated correlation matrix for endpoints								
	1	2	3	4	5	6	7	8
1 Locom Act	1							
2 Pina Reflex	0.02	1						
3 Toe Pinch	-0.02	-0.03	1					
4 Vertical Hind	0.01	0.01	-0.05	1				
5 Grip Strength	0.05	-0.03	0.08	-0.02	1			
6 Pupil Size	0.07	0.04	-0.05	0.10	-0.07	1		
7 Temperature	-0.06	-0.07	0.05	-0.13	0.04	-0.31	1	
8 Vocalization	-0.12	-0.06	-0.05	-0.02	-0.22	-0.01	0.11	1

Represents relationship between outcomes-specific evolutions.

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Hypothesis Testing

Four types of tests are considered:

- Test whether there is a treatment effect for each response separately
- Test whether there is an overall treatment effect (over all responses)
- Test whether there is a dose effect for groups of variables (such as the sensoro-motor responses and motor-affective responses)
- Test whether there is a difference among variables (such as difference between the sensoro-motor responses pinna reflex and toe pinch)

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