

A parametric Rüger test and diagnostic trials

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joint work with

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Motivation: Trial of a medical diagnostic tool

Trial plan:

- Apply diagnostic tool to 200 patients
- Assess output by $n = 3$ readers

Each patient i is classified by each reader j :

- $y_{ij} = 1$ diseased
- $y_{ij} = 0$ healthy or

Measure of performance

- $Sensitivity_j = P(Y_{ij} = 1 | \text{diseased})$
- Estimate from m diseased patients:
$$\hat{s}_j = \frac{1}{m} \sum_{i=1}^m y_{ij}$$

The proposed testing procedure

Aim: Show that a sensitivity > 0.7 can be achieved

- Hypothesis for reader j $H_j : \text{sens}_j \leq 0.7$
- Global null hypothesis $H : H_1 \cap H_2 \cap H_3$
- Test each H_j separately at level α
- If the test is significant for 2 out of 3 readers, reject global null hypothesis

Question:

What is the type I error rate of this procedure?

It is a Rüger testing procedure

- Perform a level α test for each H_j
- Reject $H : H_1 \cap \dots \cap H_n$ if k out of n individual hypotheses H_j are rejected at level α
- A sharp upper bound for the global level is $\alpha \frac{n}{k}$ (Rüger 1978)
- Test individual hypothesis on level $\alpha \frac{k}{n}$ to maintain a global level α
- k must be prespecified
- Special case $k = 1$ gives Bonferroni test

Answer, without distributional assumptions:

Type I error rate $\leq \alpha \frac{3}{2}$

Estimates of performance asymptotical multivariate normal

Sensitivity for reader j , m diseased patients:

$$\text{Estimate: } \hat{s}_j = \frac{1}{m} \sum_{i=1}^m Y_{ij}$$

$$\text{Null hypothesis: } H_j : s_j = s_0$$

$$\text{Test statistic } x_j = (\hat{s}_j - s_0) / SE(\hat{s}_j)$$

Asymptotically, by multivariate c.l.t.

$$(X|H) \sim N_3(0, \Sigma)$$

Aim:

See if we can refine boundary of type I error rate under assumption of normality

Reject H if 2 out of 3 H_j are significant

Reject H_j if $x_j > c = \phi^{-1}(1 - \alpha)$

Probability to reject $H : H_1 \cap H_2 \cap H_3$

$$\pi = P(X_1 > c, X_2 > c) + P(X_1 > c, X_3 > c) + P(X_2 > c, X_3 > c) - 2P(X_1 > c, X_2 > c, X_3 > c)$$

Simple special cases:

X uncorrelated, $\alpha \in [0; 0.5]$: $\pi = 3\alpha^2 - 2\alpha^3 < \alpha$

Perfect correlation: $\pi = \alpha$, since all X_j identical

Perfect correlation of any pair (X_s, X_t) , $s \neq t$: $\pi = \alpha$

Equally correlated test statistics

Assumptions:

$$\alpha \in [0; 0.5]$$

Test statistics $(X|H) \sim N_3(0, \Sigma)$

Equal correlations $\text{cor}(X_i, X_j) = \rho, i \neq j$

Proposition 1

Under the global null hypothesis and the above assumptions, the probability π to reject at least two out of three one-sided hypotheses at a level α is equal or below α .

Lemma for Derivative of π

$$\begin{pmatrix} X_1 \\ X_2 \\ X_3 \end{pmatrix} \sim N\left(\mu = 0, \Sigma = \begin{pmatrix} 1 & \rho_1 & \rho_2 \\ \rho_1 & 1 & \rho_3 \\ \rho_2 & \rho_3 & 1 \end{pmatrix}\right), \alpha \in [0; 0.5]$$

then

$$\frac{\partial \pi}{\partial \rho_i} \geq 0$$

if

$$\rho_j + \rho_k \leq 1 + \rho_i$$

for (i, j, k) being any permutation of $(1, 2, 3)$

- Holds for $\rho_i = \rho_k = \rho_j \leq 1$
- So π is monotonously increasing in ρ
- $\pi = \alpha$ for $\rho = 1$
- Therefore $\pi \leq \alpha$, proof of proposition 1

Location of extrema in general case

Proposition 2

Under the global null hypothesis, all local extreme values of π in (ρ_1, ρ_2, ρ_3) are on the boundary of the parameter space.

Proof: A local extreme value not on the boundary means $\frac{\partial \pi}{\partial \rho_i} = 0, i = 1, 2, 3$. Using Lemma this means

$$\begin{aligned}\rho_1 + \rho_2 &= 1 + \rho_3 \\ \rho_1 + \rho_3 &= 1 + \rho_2 \\ \rho_3 + \rho_2 &= 1 + \rho_1\end{aligned}$$

Unique solution $(\rho_1, \rho_2, \rho_3) = (1, 1, 1)$ on boundary

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Conclusions form theoretical part

- In the case of multivariate normal statistics with equal correlations, using the 2 out of 3 Rüger test with individual level α is conservative at global level α .
- For general correlation structures, the maximum of the probability to falsely reject H is found on the boundary of the correlation parameter space.
- General case is studied by considering the boundary.

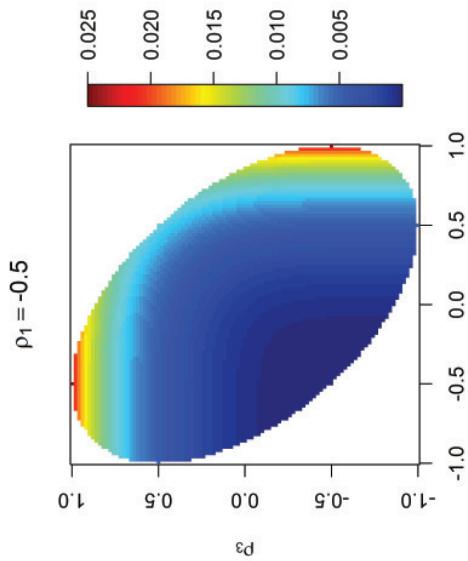
Numeric solution for general correlation structure

- Calculate probability to falsely reject H numerically (R package mvtnorm and TVPACK algorithm)
- One-sided level $\alpha = 0.025$
- Correlation matrices $\Sigma = \begin{pmatrix} 1 & \rho_1 & \rho_2 \\ \rho_1 & 1 & \rho_3 \\ \rho_2 & \rho_3 & 1 \end{pmatrix}$, n.n.d.
- with $\rho_{1,2,3} \in \{-1, -0.98, -0.96, \dots, 0.96, 0.98, 1\}$

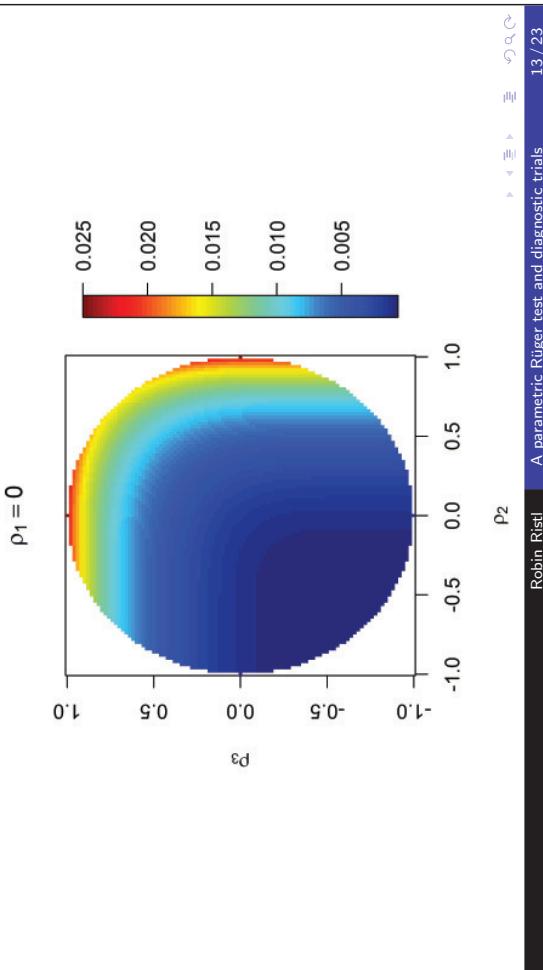
Numeric answer

$$\max(\pi) = 0.025, \text{ found for any } \rho_i = 1$$

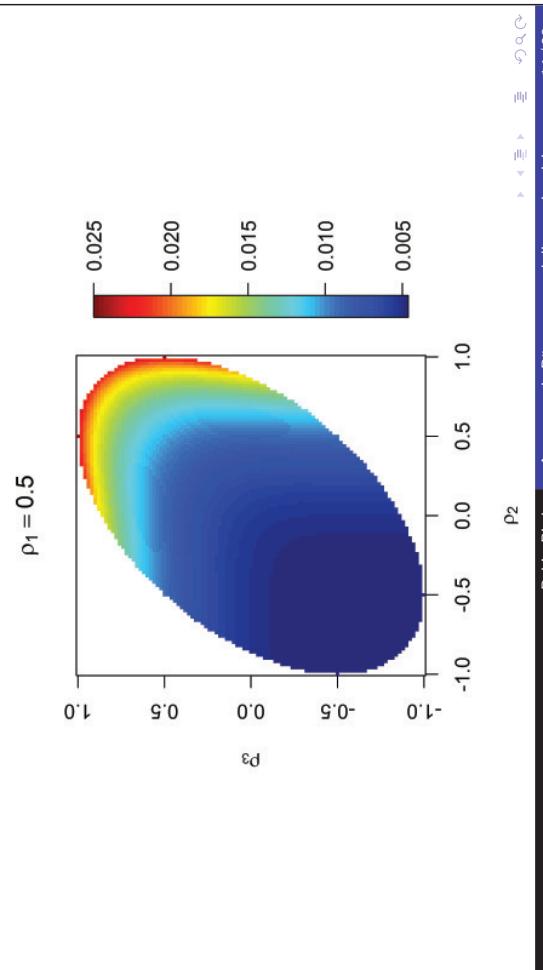
Probability to reject H - Numeric results for $\alpha = 0.025$ (1)



Probability to reject H - Numeric results for $\alpha = 0.025$ (2)

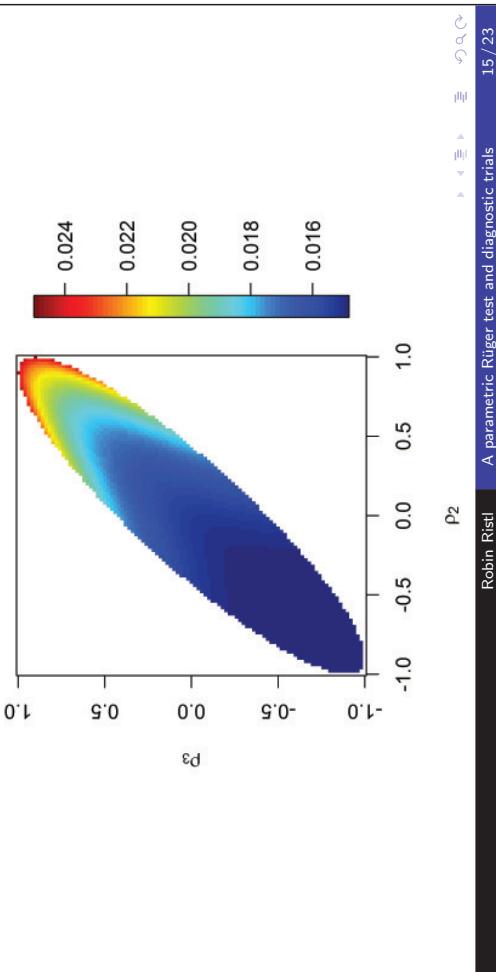


Probability to reject H - Numeric results for $\alpha = 0.025$ (3)



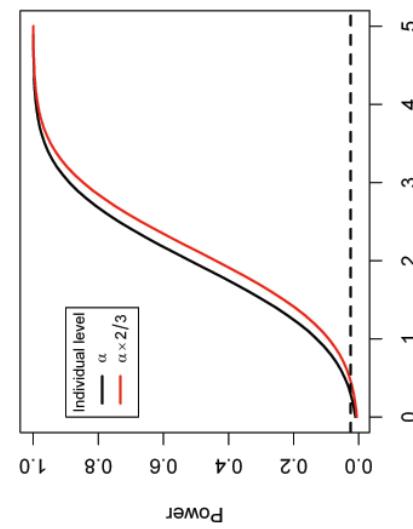
Probability to reject H_0 - Numeric results for $\alpha = 0.025$ (4)

$\rho_1 = 0.9$



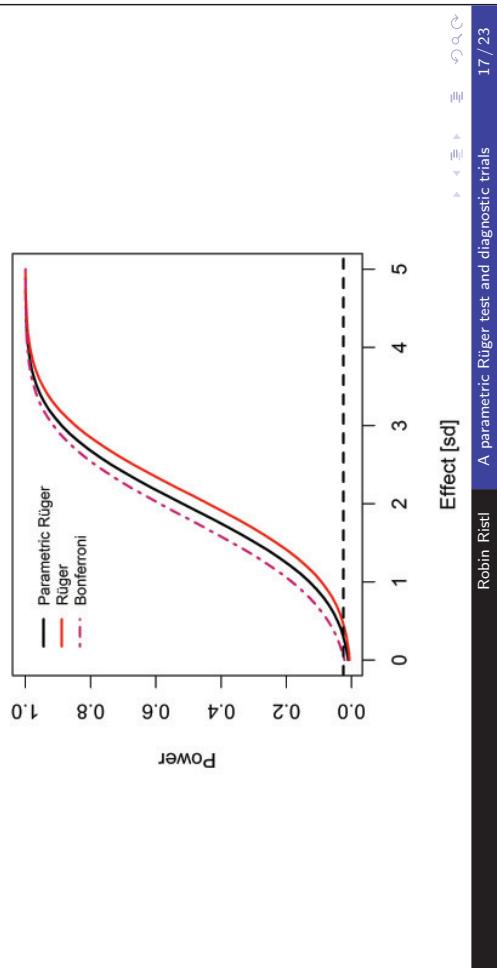
Power, equal correlations $\rho = 0.5$, all H_i false

corr = (0.5,0.5,0.5) alpha = 0.025



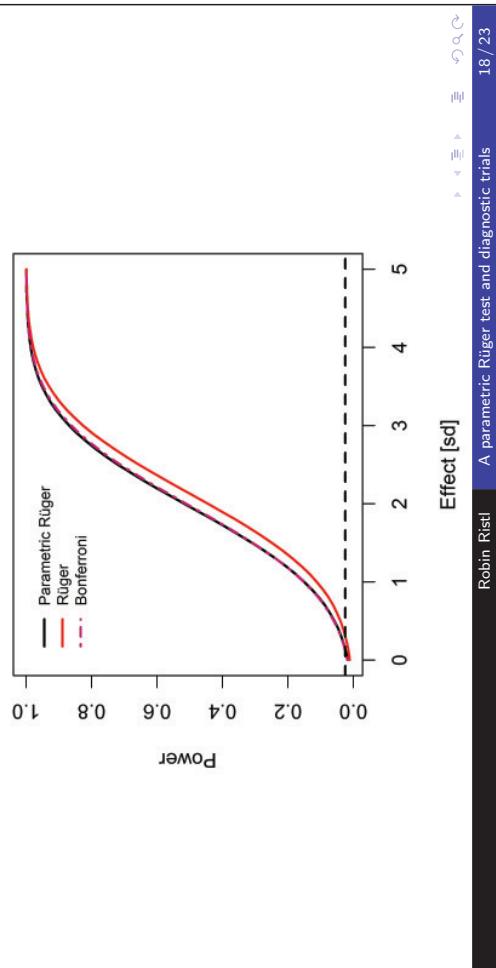
Comparing with Bonferroni test (1)

corr = (0.5,0.5,0.5) alpha = 0.025



Comparing with Bonferroni test (2)

corr = (0.75,0.75,0.75) alpha = 0.025



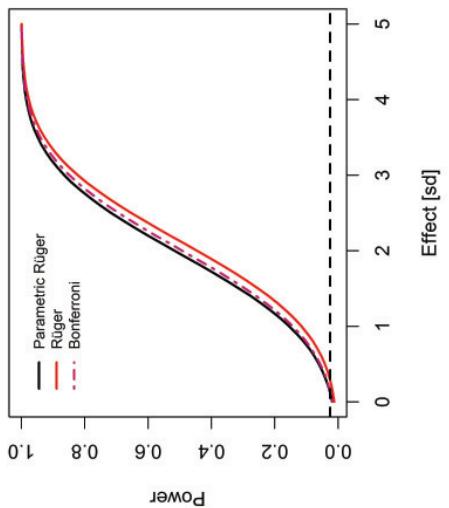
Comparing with Bonferroni test (3)

corr = (0.8,0.8,0.8) alpha = 0.025

Parametric Rüger test is superior to Bonferroni for high correlations and common effect only.

Note: Bonferroni reaches maximal type I error rate very close to α for

$$\rho_1 = \rho_2 = \rho_3 = -0.5$$



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- Several readers typical for diagnostic trial
- Specific scenario for three readers
- Rüger test free of distributional assumptions:
 $\text{Type I error rate} \leq \alpha_{\frac{3}{2}}^3$
- May assume multivariate normality
- Result for this parametric Rüger test:
 $\text{Type I error rate} \leq \alpha$
- Parametric test has higher power

Discussion

Literature

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Committee for medicinal products for human use, European Medicines Agency (2009), "Appendix 1 to the guideline on clinical evaluation of diagnostic agents (CPMP/EWP/1119/98 REV.1) on imaging agents," Doc. Ref. EMEA/CHMP/EWP/321180/2008.

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Hommel, G., Bretz, F. and Maurer W. (2011), "Multiple hypotheses testing based on ordered P values—A historical survey with applications to medical research", *Journal of Biopharmaceutical Statistics*, 21, 595-609.

Plackett, R. L. (1954), "A reduction formula for normal multivariate integrals", *Biometrika*, 41, 351-360.

Rüger, B. (1978), "Das maximale Signifikanzniveau des Tests: „Lehne H_0 ab, wenn k unter n gegebenen Tests zur Ablehnung führen“, *Metrika*, 25, 171-178.

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Thank you for your attention.

Outline for proof of Lemma

Plackett (1954) has shown $\frac{\partial \phi}{\partial \text{cov}(x_i, x_j)} = \frac{\partial^2 \phi}{\partial x_i \partial x_j}$

Applying to p gives, e.g.

$$\frac{\partial p}{\partial \rho_1} = \frac{\partial^2 p}{\partial x_1 \partial x_2} =$$

$$\int_{-\infty}^{\infty} \phi(z, z, x_3) dx_3 - 2 \int_z^{\infty} \phi(z, z, x_3) dx_3$$

$$\frac{\partial p}{\partial \rho_1} > 0 \text{ if } z > E(x_3 | x_1 = z, x_2 = z)$$

For $z > 0$, this is equivalent to $\rho_j + \rho_k < 1 + \rho_i$